RESEARCH



Association between bone mineral density and scoliosis: a two-sample mendelian randomization study in european populations



Fangjun Yang¹ and Jiantao Wen^{2*}

Abstract

Background Previous studies have shown that bone mineral density (BMD) has a certain impact on scoliosis. However, up to now, there is no clear evidence that there is a causal association between the two. The aim of this study is to investigate whether there is a causal association between BMD at different body positions and scoliosis by two-sample Mendelian randomization (MR).

Methods Genetic variants (SNPS) strongly associated with BMD (total body BMD (TB-BMD), lumbar spine BMD (LS-BMD), femoral neck BMD (FN-BMD), heel BMD (HE-BMD), and forearm BMD (FA-BMD)) were extracted from GEFOS and genome-wide association analysis (GWAS) databases SNPs) were used as instrumental variables (IVs). Scoliosis was also selected from the Finnish database as the outcome. Inverse variance weighting (IVW) method was used as the main analysis method, and multiple sensitivity analysis was performed by combining weighted median, MR-Egger, MR Multi-effect residuals and outliers.

Results IVW results showed that TB-BMD (OR = 0.83, 95%CI: 0.66–1.55 P = 0.13), LS-BMD (OR = 0.72, 95%CI: 0.52–0.99, P = 0.04), FN-BMD (OR = 0.74, 95%CI: 0.50–1.09, P = 0.13), FA-BMD (OR = 0.95,95%CI: 0.70–1.28, P = 0.75), HE-BMD (OR = 0.91, 95%CI: 0.77–1.08, P = 0.29). Sensitivity analyses showed no evidence of pleiotropy or heterogeneity (p > 0.05) (MR-PRESSO and Cochrane). The results were further validated by leave-one-out test and MR-Egger intercept, which confirmed the robustness of the study results.

Conclusion In conclusion, the present study demonstrates that the causal role of genetic prediction of scoliosis increases with decreasing lumbar BMD. There was no evidence that BMD at the remaining sites has a significant causal effect on scoliosis. Our results suggest that the lumbar spine BMD should be routinely measured in the population at high risk of scoliosis. If osteoporosis occurs, appropriate treatment should be given to reduce the incidence of scoliosis.

Clinical trial number Not applicable

Keywords Mendelian randomization, Scoliosis, Bone mineral density, Causality

*Correspondence: Jiantao Wen 465545037@qq.com ¹Department of orthopedic, Gansu Provincial Hospital of Traditional Chinese Medicine, Lanzhou, China ²Department of Pediatric Spine Surgery, Gansu Provincial Hospital of Traditional Chinese Medicine, Lanzhou, China



© The Author(s) 2024. **Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicate of the original autory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/. The Creative Commons Dedication waiver (http://creativecommons.org/publicdomain/zero/1.0/) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

Introduction

Scoliosis is a complex three dimensional (3D) structural deformity characterized by more than 10° of scoliosis on coronal radiographs of the spine with axial rotation and sagittal deviation [1]. According to the presence or absence of a clear cause, it can be divided into idiopathic and non-idiopathic [2, 3]. The global prevalence ranges from 0.47–5.2% [4, 5]. Despite extensive experimental and clinical studies over the years, its etiology and pathogenesis remain poorly understood. The etiology theory involves genetics, biomechanics, nervous system, endocrine system, spinal cord growth and bone metabolism [6]. Previous studies [7-11] have suggested that decreased bone mineral density (BMD) and even osteoporosis play an important role in the pathogenesis of scoliosis, but these studies have not suggested a potential cause-and-effect relationship.

To determine the causal relationship between BMD and scoliosis, we conducted a two-sample Mendelian randomization (MR) Study using public genome-wide Association Study (GWAS) data to explore the association. Mendelian randomization (MR) studies refer to the use of genetic variants (single nucleotide polymorphisms, SNPs) as instrumental variables (IVs) to infer causal associations of exposure levels (such as biomarkers) with outcomes [12]. Because SNPS are randomly assigned at conception (according to Mendelian's second law), environmental confounders and disease states that develop later in life do not affect germline genetic susceptibility. Therefore, MR Limits bias due to confounding and reverse causality and allows causal inference [13, 14].

Methods

Study design

The flow of the study design is illustrated in Fig. 1. The IVs required for the MR Analysis must satisfy the following three assumptions [15-17]. (1) The IVs used were strongly associated with exposure (BMD at different sites); (2) The selected IVs were not related to potential confounders; (3) IVs could affect the risk of outcome (scoliosis) only through exposure.

Data sources

Exposure data included total body BMD (TB-BMD), lumbar spine BMD (LS-BMD), femoral neck BMD (FN-BMD), forearm BMD (FA-BMD), and heel BMD (HE-BMD). Above all data from osteoporosis GEFOS genetic factors alliance (http://www.gefos.org/) and IEU Open GWAS (https://gwas.mrcieu.ac.uk/datasets/) summary statistics.

Scoliosis genetic statistics data from 2021 released R5 FinnGen alliance (https://r5.risteys.finngen.fi/). The large



SNP, single nucleotide polymorphism, TB total body, LS lumbar spine,

FN femoral neck, HE heels, FA forearm, BMD, bone mineral density

GWAS in Finns contained 1168 cases and 164,682 controls, which yielded 16,380,270 SNPs for analysis after adjusting for factors such as age, sex, and genotyping batch [18]. Ethical approval and informed consent were provided for the above data, and the writing process followed the requirements of STROBE-MR, the reporting code for MR Studies. Data details are provided in Table 1.

Selection of IVs

The choice of IVs should conform to the above three assumptions. The SNPs ($P < 5 \times 10^{-8}$) that were closely associated with BMD at the five sites were screened from the corresponding databases (Assumption 1). Second, SNPs pairwise correlation coefficients in linkage disequilibrium (LD) must satisfy $r^2 < 0.001$ and kb = 10,000to be considered independent (Assumption 2). In order to ensure that gene variants independently of potential confounders (Assumption 3), using PhenoScanner (htt p://www.phenoscanner.medschl.cam.ac.uk/) to screen and remove possible confounding factors. Known confounders that may be associated with the outcome (scoliosis) include vitamin D [23], age at menarche [24], body weight [25], BMI [26, 27], leptin [28], estrogen [29], melatonin [30], etc. The strength of instrumental variables was estimated using the F-statistic, which was calculated as $F = (R^2/k)/([1-R^2]/[n-k-1])$. R² represents the proportion of variance explained by the exposed SNP tools, k represents the number of tools, and N represents the exposed sample size [31, 32]. F-statistics of less than 10 were considered to indicate weak instrumental variable bias and were excluded [33]. These SNPs were then used as IVs to assess causality between exposure and outcome in the MR Analysis.

MR analysis

In this two-sample MR Design, we used the following five methods to analyze the causal association between BMD of different body parts and scoliosis separately [34–36]. These methods include inverse variance weighting (IVW), weighted median (WM), simple median (SM), weighted median estimator (WME) and MR-Egger regression. Among them, IVW calculates the odds ratio (OR) as the primary method, which is typical and routine in MR, and the slope of the weighted regression of the outcome effect on the exposure effect represents the

 Table 1
 Bone mineral density (BMD) GWAS data summary

outcome estimate (with an intercept restricted to zero). In addition, the other four methods were used as tests for the robustness of the primary outcome. Heterogeneity of individual estimates of genetic variation was assessed using Cochran's Q test. If Cochran's Q P>0.05 and there was no evidence of heterogeneity, the fixed effect IVW method was used. If there was significant heterogeneity (P<0.05 by Cochran's Q test), the random effects IVW method was used [37–39].

Sensitivity analysis

To ensure that IVs was independent of outcomes other than exposure, we used different approaches to exclude potential effects. Firstly, Pleiotropy residual sum and outlier (MR-PRESSO) was applied to test and calibrate the outliers of horizontal pleiotropy, and the outliers in IVs were removed [34]. We used MR-Egger regression to account for horizontal pleiotropy, with a *P* value of more than 0.05 for the intercept indicating the absence of horizontal pleiotropy [40]. Horizontal pleiotropy was tested by drawing a funnel plot. If the funnel plot shows a symmetric shape, this usually means that there is no obvious pleiotropy.

Leave-one-out sensitivity test is to observe whether the results will change significantly after removing a specific SNP. If the results remain relatively stable after removing any SNP one by one, it indicates that the overall error line will change within a small range after removing any SNP. Specifically, each SNP was removed one by one, and then the meta-effects of the remaining SNPs were calculated, in which the influence of each SNP was separately evaluated by IVW analysis and represented by forest plot.

Analysis software

Statistical analysis of all data was performed using R Studio (version 4.3.1), two-sample MR (version 0.5.7) and MR-PRESSO (version 1.0) software packages [41]. Results are presented as odds ratios (ORs) with 95% confidence intervals (95% CI), and if p<0.05 was considered statistically significant.

Results

Overall, we obtained SNPs that met the three assumptions of MR Analysis and could be used for MR Analysis (Supplementary Table 1). In the heterogeneity test, all

Table T bone mineral density (bmb) dwas data summary						
GWAS ID	Trait	Consortium/	Sample size	Trait	Reference	
ebi-a-GCST00 5348	TB-BMD	GWAS meta- analysis	56 284	16 162 733	Medina et al., 2018 [19]	
ieu-a-982	LS-BMD	GEFOS	28 498	10 582 867	Zheng et al., 2015 [20]	
ieu-a-980	FN-BMD	GEFOS	32 735	10 586 900	Zheng et al., 2015 [20]	
ieu-a-977	FA-BMD	GEFOS	8 143	9 955 366	Surakka et al., 2020 [21]	
ebi-a-GCST00 6979	Heel-BMD	GEFOS	426 824	13 705 641	Morris et al., 2019 [22]	

		exposures				
		TB-BMD	LS-BMD	FN-BMD	FA-BMD	HE-BMD
SNPs, n		63	18	14	11	419
IVW	OR	0.83	0.72	0.74	0.95	0.91
	95%CI	0.66-1.55	0.52-0.99	0.50-1.09	0.70-1.28	0.77-1.08
	<i>P</i> value	0.13	0.04	0.13	0.75	0.29
Weighted median	OR	0.76	0.57	0.78	1.22	1.01
	95%CI	0.54-1.07	0.36-0.90	0.47-1.31	0.82-1.83	0.76-1.37
	Pvalue	0.12	0.01	0.36	0.32	0.90
Weighted mode	OR	0.60	0.52	0.78	1.26	1.11
	95%CI	0.33-1.10	0.23-1.20	0.37-1.67	0.80-1.96	0.81-1.51
	Pvalue	0.10	0.14	0.53	0.33	0.51
MR Egger	OR	1.43	1.17	2.02	1.22	1.06
	95%CI	0.80-2.58	0.30-4.56	0.29-14.00	0.51-2.91	0.80-1.42
	<i>P</i> value	0.24	0.82	0.48	0.66	0.65
Simple mode	OR	0.57	0.54	0.78	0.95	0.74
	95%CI	0.26-1.25	0.22-1.33	0.34-1.73	0.40-2.20	0.38–1.46
	Pvalue	0.17	0.20	0.54	0.90	0.39

Table 2 Results of mendelian randomization analysis

The boldface values represent the number of SNPS, the results of the main research methods, and the direct causal relationship (P<0.05)



Fig. 2 Analysis results of leave-one method

IVs did not show significant heterogeneity (p>0.05), so a fixed effects model was used for the five-item MR Analysis when calculating the IVW.

IVW results were TB-BMD (OR=0.83, 95%CI: 0.66–1.55, P=0.13); LS-BMD (OR=0.72, 95%CI: 0.52–0.99, P=0.04) FN-BMD (OR=0.74, 95%CI: 0.50–1.09, P=0.13); FA-BMD (OR=0.95,95%CI: 0.70–1.28, P=0.75); HE-BMD (OR=0.91, 95%CI: 0.77–1.08, P=0.29). The complete results of the five MR Analyses are shown in Table 2. MR Study found a negative causal relationship between lumbar BMD and scoliosis (LS-BMD IVW OR=0.72, 95%CI: 0.52–0.99, P=0.04). There was no causal association between BMD at the remaining sites and scoliosis.

In the five MR Analyses performed, the *P* value for the intercept in the MR-Egger regression was higher than 0.05 (Table 2). No outliers were found by MR-PRESSO test. MR-Egger intercept and MR-PRESSO analysis

showed that there was no pleiotropy of the above exposures on scoliosis, and there was no heterogeneity between them. In addition, the results of leave-one-out analyses performed and plotted show that no single SNP had an effect on the overall causal estimate (Fig. 2).

To further demonstrate the credibility of our findings, we plotted funnel plots and scatter plots for visual assessment of horizontal pleiotropy and heterogeneity. The distribution of causal effects shown in the funnel plot was basically symmetric and no obvious bias was observed (Fig. 3), and we found possible outliers in the IV of FN-BMD and FA-BMD in the scatter plot (Fig. 4). However, MR-PRESSO analysis showed no significant outliers (global test P > 0.05). Thus, the relationship between BMD and scoliosis provides insufficient support for horizontal pleiotropy. Reverse MR Analysis showed no causal relationship between scoliosis and BMD at the five sites (Supplementary Table 2).



Fig. 3 Funnel plot of causality between BMD and scoliosis



Fig. 4 Scatter plot of causality between BMD and scoliosis

Discussion

So far, A variety of imaging methods have been used to measure BMD, such as dual-energy X-ray absorptiometry (DXA), quantitative ultrasound system (QUS), computed tomography, and quality control techniques (QCT) [42]. Many studies [43–46] have demonstrated the fact that BMD is reduced in patients with scoliosis by the above methods, but these studies have not proved whether there is a causal association between BMD at different locations and the presence or absence of scoliosis.

To the best of our knowledge, this is the first study to investigate the causal association between bone mineral density at different locations and scoliosis using twosample MR Analysis. Studies have shown that lower lumbar BMD is associated with a significantly increased risk of scoliosis in a European population. BMD at other locations was not causally associated with scoliosis.

Cheng et al. [47] demonstrated that low LS-BMD may play an important role in the etiology and pathogenesis of scoliosis through bone histological studies, and low BMD is caused by metabolic disorders leading to a reduction in the number of osteoclasts in the trabecular compartment. Through DNA analysis of 198 girls diagnosed with AIS, Suh et al. [48] found that vitamin D receptor (VDR) BsmI polymorphism was associated with LS-BMD in AIS girls, and low bone mass may affect abnormal spinal growth patterns through VDR gene BsmI locus polymorphism.

The above studies have proved that LS-BMD plays an important role in scoliosis through various ways from different perspectives, but they have not revealed the direct causal relationship between the two determined by genes. Our results have proved the above conclusions, and our results are not affected by confounding factors, so the results are more reliable.

Handa et al. [49] found that in a longitudinal observation study in mice, the growth plate was thickened and osteoblasts were reduced over time, suggesting that impaired endochondral ossification was the cause of scoliosis. Reduction of bone mineral density and degradation of bone microstructure were also observed. This suggests that defects in endochondral ossification may impair growth, leading to scoliosis and decreased BMD. Therefore, more studies are needed in the future.

Our study has several strengths. Firstly, we applied the MR Method for the first time to investigate the causal relationship between BMD at different bone sites and scoliosis, avoiding potential confounding factors and reverse causality. Second, our data were derived from the GWAS, FinnGen, and the GEFOS consortium summary data, and the results were consistent across different datasets, ensuring the reliability of our findings.

However, this study still has some potential limitations. First, the database used in this study was based on populations of European ancestry, and it is unclear whether the results would apply to populations of non-European ancestry. Second, BMD and prevalence of scoliosis vary by age and sex, but this study was analyzed based on data from GWAS pooled levels, which does not allow subgroups to assess effects according to different age and sex. Third, the results of this study showed a significant causal relationship between LS-BMD and scoliosis, but not TB-BMD, FN-BMD, HE-BMD or FA-BMD. Further MR Studies with larger sample sizes or randomized controlled trials are needed to confirm these findings. Finally, Unobserved pleiotropy is a major limitation of MR Studies, which may influence conclusions that assess the association between genetically predicted BMD and scoliosis risk.

Conclusions

In conclusion, this bi-directional two-sample MR Study found a causal relationship between LS-BMD and scoliosis, but found no evidence of a causal relationship between BMD at other sites and scoliosis. The lower the LS-BMD level, the higher the risk of scoliosis.

Abbreviations

BMD	Bone mineral density
DXA	Dual energy X-ray absorptiometry
QUS	Quantitative ultrasonography
CT	Computerized tomography
GWAS	Genome Wide Association Study
MR	Mendelian randomization
SNPs	Single nucleotide polymorphisms
IVs	Instrumental variables
ТВ	Total body
LS	Lumbar spine
FN	Femoral neck
HE	Heels
FA	Forearm
IVW	Inverse variance weighting
WM	Weighted median
SM	Simple median
WME	Weighted median estimator

Supplementary Information

The online version contains supplementary material available at https://doi.or g/10.1186/s41065-024-00352-w.

Supp	lementary	Material	1

Supplementary Material 2

Acknowledgements

The authors are grateful to the participants of all the GWASs used in this manuscript and the investigators who made these GWAS data publicly available.

Author contributions

Fangjun Yang and Jiantao Wen conceived and designed this study. Fangjun Yang conducted the data acquisition, analyzed and interpreted the data, and drafted the manuscript. Fangjun Yang contributed to the analysis and interpretation of data, proofread the data, and critically revised the manuscript. All authors approved the final submitted version.

Funding

 Major Project of Gansu Province Joint Scientific Research Fund (23JRRA1529): Research on prevention and treatment of adolescent idiopathic scoliosis with integrated Chinese and Western medicine;
 Natural Science Foundation of Gansu Province (22JR5 RA632): Screening study of protein molecular markers associated with adolescent idiopathic scoliosis in serum of adolescent spine health cohort.

Data availability

No datasets were generated or analysed during the current study.

Declarations

Ethics approval and consent to participate

Not applicable.

Informed consent statement

Not applicable.

Competing interests

The authors declare no competing interests.

Received: 6 September 2024 / Accepted: 19 November 2024 Published online: 31 December 2024

References

- Burwell RG, Clark EM, Dangerfield PH, et al. Adolescent idiopathic scoliosis (AIS): a multifactorial cascade concept for pathogenesis and embryonic origin[J]. Scoliosis Spinal Disord. 2016;11:8. https://doi.org/10.1186/s13013-01 6-0063-1
- Donzelli S, Zaina F, Lusini M, et al. In favour of the definition adolescents with idiopathic scoliosis: juvenile and adolescent idiopathic scoliosis braced after ten years of age, do not show different end results. SOSORT Award Win 2014[J] Scoliosis. 2014;9:7. https://doi.org/10.1186/1748-7161-9-7
- Takahashi Y, Kou I, Takahashi A, et al. A genome-wide association study identifies common variants near LBX1 associated with adolescent idiopathic scoliosis[J]. Nat Genet. 2011;43(12):1237–40. https://doi.org/10.1038/ng.974
- Konieczny MR, Senyurt H, Krauspe R. Epidemiology of adolescent idiopathic scoliosis[J]. J Child Orthop. 2013;7(1):3–9. https://doi.org/10.1007/s11832-01 2-0457-4
- Peng Y, Wang SR, Qiu GX, et al. Research progress on the etiology and pathogenesis of adolescent idiopathic scoliosis[J]. Chin Med J (Engl). 2020;133(4):483–93. https://doi.org/10.1097/cm9.000000000000652
- Aulia TN, Djufri D, Gatam L, et al. Etiopathogenesis of adolescent idiopathic scoliosis (AIS): role of genetic and environmental factors[J]. Narra J. 2023;3(3):e217. https://doi.org/10.52225/narra.v3i3.217
- Daffner SD, Vaccaro AR. Adult degenerative lumbar scoliosis[J]. Am J Orthop (Belle Mead NJ). 2003;32(2):77–82. discussion 82.
- Velis KP, Healey JH, Schneider R. Osteoporosis in unstable adult scoliosis[J]. Clin Orthop Relat Res, 1988;11(237):132–41.
- Carter OD, Haynes SG. Prevalence rates for scoliosis in US adults: results from the first National Health and Nutrition Examination Survey[J]. Int J Epidemiol. 1987;16(4):537–44. https://doi.org/10.1093/ije/16.4.537
- Ding WY, Yang DL, Cao LZ, et al. Intervertebral disc degeneration and bone density in degenerative lumbar scoliosis: a comparative study between patients with degenerative lumbar scoliosis and patients with lumbar stenosis[J]. Chin Med J (Engl). 2011;124(23):3875–8.
- Rubin J, Cleveland RJ, Padovano A, et al. Lumbar scoliosis in Postmenopausal Women increases with age but is not Associated with Osteoporosis[J]. J Endocr Soc. 2021;5(5):bvab018. https://doi.org/10.1210/jendso/bvab018

- 12. Zheng J, Baird D, Borges MC, et al. Recent developments in mendelian randomization Studies[J]. Curr Epidemiol Rep. 2017;4(4):330–45. https://doi.org/10.1007/s40471-017-0128-6
- Lawlor DA, Harbord RM, Sterne JA, et al. Mendelian randomization: using genes as instruments for making causal inferences in epidemiology[J]. Stat Med. 2008;27(8):1133–63. https://doi.org/10.1002/sim.3034
- 14. Emdin CA, Khera AV, Kathiresan S. Mendelian Randomization[J] Jama. 2017;318(19):1925–6. https://doi.org/10.1001/jama.2017.17219
- Lawlor DA, Commentary. Two-sample mendelian randomization: opportunities and challenges[J]. Int J Epidemiol. 2016;45(3):908–15. https://doi.org/10.1 093/ije/dyw127
- Davey Smith G, Hemani G. Mendelian randomization: genetic anchors for causal inference in epidemiological studies[J]. Hum Mol Genet. 2014;23(R1):R89–98. https://doi.org/10.1093/hmg/ddu328
- Birney E. Mendelian Randomization[J]. Cold Spring Harb Perspect Med. 2022;12(4). https://doi.org/10.1101/cshperspect.a041302
- Lai B, Jiang H, Gao Y, et al. Causal effects of gut microbiota on scoliosis: a bidirectional two-sample mendelian randomization study[J]. Heliyon. 2023;9(11):e21654. https://doi.org/10.1016/j.heliyon.2023.e21654
- Medina-Gomez C, Kemp JP, Trajanoska K, et al. Life-Course Genome-Wide Association Study Meta-Analysis of Total Body BMD and Assessment of Age-Specific Effects[J]. Am J Hum Genet. 2018;102(1):88–102. https://doi.org/10.1 016/j.ajhg.2017.12.005
- 20. Zheng HF, Forgetta V, Hsu YH, et al. Whole-genome sequencing identifies EN1 as a determinant of bone density and fracture[J]. Nature. 2015;526(7571):112–7. https://doi.org/10.1038/nature14878
- Surakka I, Fritsche LG, Zhou W, et al. MEPE loss-of-function variant associates with decreased bone mineral density and increased fracture risk[J]. Nat Commun. 2020;11(1):4093. https://doi.org/10.1038/s41467-020-17315-0
- 22. Morris JA, Kemp JP, Youlten SE, et al. An atlas of genetic influences on osteoporosis in humans and mice[J]. Nat Genet. 2019;51(2):258–66. https://doi.org /10.1038/s41588-018-0302-x
- 23. Balioglu MB, Aydin C, Kargin D, et al. Vitamin-D measurement in patients with adolescent idiopathic scoliosis[J]. J Pediatr Orthop B. 2017;26(1):48–52. https://doi.org/10.1097/bpb.0000000000320
- Chang HK, Chang DG, Myong JP, et al. Bone mineral density among Korean females aged 20–50 years: influence of age at menarche (the Korea National Health and Nutrition Examination Survey 2008–2011)[J]. Osteoporos Int. 2017;28(7):2129–36. https://doi.org/10.1007/s00198-017-3997-0
- Clark EM, Taylor HJ, Harding I, et al. Association between components of body composition and scoliosis: a prospective cohort study reporting differences identifiable before the onset of scoliosis[J]. J Bone Min Res. 2014;29(8):1729– 36. https://doi.org/10.1002/jbmr.2207
- Jeon K, Kim DI. The Association between Low Body Weight and Scoliosis among Korean Elementary School Students[J]. Int J Environ Res Public Health. 2018;15(12). https://doi.org/10.3390/ijerph15122613
- Kim S, Uhm JY, Chae DH, et al. Asian Nurs Res (Korean Soc Nurs Sci).
 2020;14(1):24–9. https://doi.org/10.1016/j.anr.2019.12.003. Low Body Mass Index for Early Screening of Adolescent Idiopathic Scoliosis: A Comparison Based on Standardized Body Mass Index Classifications[J].
- Tam EMS, Liu Z, Lam TP, et al. Lower muscle Mass and Body Fat in adolescent idiopathic scoliosis are Associated with abnormal leptin Bioavailability[J]. Spine (Phila Pa 1976). 2016;41(11):940–6. https://doi.org/10.1097/brs.000000 000001376
- Warren MP, Brooks-Gunn J, Hamilton LH, et al. Scoliosis and fractures in young ballet dancers. Relation to delayed menarche and secondary amenorrhea[J]. N Engl J Med. 1986;314(21):1348–53. https://doi.org/10.1056/nejm198605223 142104
- Golan D, Halhal B, Glass-Marmor L, et al. Vitamin D supplementation for patients with multiple sclerosis treated with interferon-beta: a randomized controlled trial assessing the effect on flu-like symptoms and immunomodulatory properties[J]. BMC Neurol. 2013;13:60. https://doi.org/10.1186/1471-23 77-13-60
- 31. Palmer TM, Lawlor DA, Harbord RM, et al. Using multiple genetic variants as instrumental variables for modifiable risk factors[J]. Stat Methods Med Res. 2012;21(3):223–42. https://doi.org/10.1177/0962280210394459
- Park JH, Wacholder S, Gail MH, et al. Estimation of effect size distribution from genome-wide association studies and implications for future discoveries[J]. Nat Genet. 2010;42(7):570–5. https://doi.org/10.1038/ng.610

- Bowden J, Del Greco MF, Minelli C, et al. Assessing the suitability of summary data for two-sample mendelian randomization analyses using MR-Egger regression: the role of the l2 statistic[J]. Int J Epidemiol. 2016;45(6):1961–74. https://doi.org/10.1093/ije/dyw220
- Verbanck M, Chen CY, Neale B, et al. Detection of widespread horizontal pleiotropy in causal relationships inferred from mendelian randomization between complex traits and diseases[J]. Nat Genet. 2018;50(5):693–8. https:// doi.org/10.1038/s41588-018-0099-7
- Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression[J]. Int J Epidemiol. 2015;44(2):512–25. https://doi.org/10.1093/ije/d yv080
- Wang W, Tan JS, Hua L, et al. Genetically predicted obesity causally increased the risk of Hypertension disorders in Pregnancy[J]. Front Cardiovasc Med. 2022;9:888982. https://doi.org/10.3389/fcvm.2022.888982
- Papadimitriou N, Dimou N, Tsilidis KK, et al. Physical activity and risks of breast and colorectal cancer: a mendelian randomisation analysis[J]. Nat Commun. 2020;11(1):597. https://doi.org/10.1038/s41467-020-14389-8
- Shi Q, Wang Q, Wang Z, et al. Systemic inflammatory regulators and proliferative diabetic retinopathy: a bidirectional mendelian randomization study[J]. Front Immunol. 2023;14:1088778. https://doi.org/10.3389/fimmu.2023.10887 78
- Yuan S, Kim JH, Xu P, et al. Causal association between celiac disease and inflammatory bowel disease: a two-sample bidirectional mendelian randomization study[J]. Front Immunol. 2022;13:1057253. https://doi.org/10.3389/fi mmu.2022.1057253
- Burgess S, Thompson SG. Interpreting findings from mendelian randomization using the MR-Egger method[J]. Eur J Epidemiol. 2017;32(5):377–89. https://doi.org/10.1007/s10654-017-0255-x
- Hemani G, Zheng J, Elsworth B, et al. The MR-Base platform supports systematic causal inference across the human phenome[J]. Elife. 2018;7. https://doi. org/10.7554/eLife.34408
- 42. Thomas KA, Cook SD, Skalley TC, et al. Lumbar spine and femoral neck bone mineral density in idiopathic scoliosis: a follow-up study[J]. J Pediatr Orthop. 1992;12(2):235–40. https://doi.org/10.1097/01241398-199203000-00016
- Cheng JC, Hung VW, Lee WT, et al. Persistent osteopenia in adolescent idiopathic scoliosis–longitudinal monitoring of bone mineral density until skeletal maturity[J]. Stud Health Technol Inf. 2006;123:47–51.
- 44. Pourabbas Tahvildari B, Erfani MA, Nouraei H, et al. Evaluation of bone mineral status in adolescent idiopathic scoliosis[J]. Clin Orthop Surg. 2014;6(2):180–4. https://doi.org/10.4055/cios.2014.6.2.180
- Sarioglu O, Gezer S, Sarioglu FC, et al. Evaluation of vertebral bone mineral density in scoliosis by using quantitative computed tomography[J]. Pol J Radiol. 2019;84:e131–5. https://doi.org/10.5114/pjr.2019.84060
- 46. Cheuk KY, Hu Y, Tam EMS, et al. Bone measurements at multiple skeletal sites in adolescent idiopathic scoliosis-an in vivo correlation study using DXA, HRpQCT and QCT[J]. Arch Osteoporos. 2019;14(1):70. https://doi.org/10.1007/s1 1657-019-0621-2
- Cheng JC, Tang SP, Guo X, et al. Osteopenia in adolescent idiopathic scoliosis: a histomorphometric study[J]. Spine (Phila Pa 1976). 2001;26(3):E19–23. https://doi.org/10.1097/00007632-200102010-00002
- Suh KT, Eun IS, Lee JS. Polymorphism in vitamin D receptor is associated with bone mineral density in patients with adolescent idiopathic scoliosis[J]. Eur Spine J. 2010;19(9):1545–50. https://doi.org/10.1007/s00586-010-1385-y
- Handa M, Demura S, Yokogawa N, et al. Characteristics of scoliosis in mice Induced by Chondrocyte-specific inactivation of L-type amino acid transporter 1[J]. Spine (Phila Pa 1976). 2024;49(4):285–93. https://doi.org/10.1097/ brs.000000000004842

Publisher's note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.